# Sequential Decision Making Lecture 3 : Beyond Classical Bandits

Emilie Kaufmann



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## **Recap from last class**

Several important ideas to tackle the exploration/exploitation challenge in a simple multi-armed bandit model with independent arms :

- Explore then Commit
- $\triangleright$   $\varepsilon$ -greedy
- Optimistic algorithms : Upper Confidence Bounds strategies
- Bayesian algorithms : Thompson Sampling

Some of these can be extended to more realistic **structured** models that are suited for different applications.

## Outline

#### 1 Contextual Bandits

- Solving Linear Bandits
   Lin-UCB
  - Linear Thompson Sampling
- 3 Other variants of the classical MAB
- **4** Beyond maximizing rewards

## **Contextual Bandits**

#### Example : movie recommendation



What movie should Netflix recommend to a particular user, given the ratings provided by previous users?

to make good recommendation, we should take into account the characteristics of the movies / users

#### Contextual bandit problem : at time t

- a context c<sub>t</sub> is observed
- $\blacktriangleright$  an arm  $A_t$  is chosen
- ▶ a reward  $R_t$  that depends on  $c_t$ ,  $A_t$  is received.

## Mixing bandits and regression models

- A contextual bandit model incorporates two components :
  - a sequential interaction protocol : pick an arm, receive a reward
  - > a regression model for the dependency between context and reward

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#### A (stochastic) contextual bandit model incorporates two components :

- a sequential interaction protocol : pick an arm, receive a (random) reward
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#### General stochastic contextual bandit model

In each round t, the agent

▶ observes a context  $c_t \in C$ 

(user characteristics)

- ▶ selects an arm  $A_t \in A_t$  (an item out of a possibly changing pool)
- the agent receives a reward

$$r_t = f_{A_t}(c_t) + \varepsilon_t$$

where  $\varepsilon_t$  is an independent noise :  $\mathbb{E}[\varepsilon_t] = 0$ .

 $f_a:\mathcal{C} o \mathbb{R}$  maps a context c to the average reward of arm  $a, f_a(c)$ 

## **Examples**

#### Example 1

- user t : descriptor  $c_t \in \mathbb{R}^p$
- item a : descriptor  $\theta_a \in \mathbb{R}^p$

$$r_t = \theta_{A_t}^\top c_t + \varepsilon_t$$

Linear function  $f_a(c) = \theta_a^\top c$ 

**Observation** : if  $\mathcal{A}_t = \{1, \dots, K\}$  is a fixed set of items

- ▶ the model is parameterized by  $\theta_1, \theta_2, \ldots, \theta_K \in (\mathbb{R}^p)^K$
- ▶ it can also be rewritten  $r_t = \theta_{\star}^{\top}(x_{t,A_t}) + \varepsilon_t$  with

$$\theta_{\star} = \begin{pmatrix} \theta_{1} \\ \cdots \\ \theta_{a} \\ \cdots \\ \theta_{K} \end{pmatrix} \in \mathbb{R}^{p \times K}, \quad x_{t,a} = \begin{pmatrix} 0 \\ \cdots \\ c_{t} \\ \cdots \\ 0 \end{pmatrix} \in \mathbb{R}^{p \times K}$$

 $x_{t,a}$  : feature vector for the user-item pair (t,a)Emilie Kaufmann | CRIStAL

## **Examples**

#### Example 2

- user t : descriptor  $c_t \in \mathbb{R}^p$
- item a : descriptor  $x_a \in \mathbb{R}^{p'}$
- → build a user-item feature vector for (t, a) :  $x_{t,a} \in \mathbb{R}^d$

(feature engineering)

$$\mathbf{r}_t = \theta_\star^\top \mathbf{x}_{t, \mathbf{A}_t} + \varepsilon_t$$

#### Observation :

▶ the model is parameterized by  $\theta_{\star} \in \mathbb{R}^d$ 

▶ in each round *t*, the user-item feature vectors belong to the set

$$\mathcal{X}_t = \{x_{t,a}, a \in \mathcal{A}_t\} \subseteq \mathbb{R}^d$$

▶ picking an arm  $a \leftrightarrow$  picking a feature vector  $x_t \in \mathcal{X}_t$  $r_t = \theta_t^\top x_t + \varepsilon_t$ 

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## **Examples**

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▶ picking an arm  $a \leftrightarrow$  picking a feature vector  $x_t \in \mathcal{X}_t$  $r_t = f_\star(x_t) + \varepsilon_t$ 

## **Two formulations**

## Contextual MAB, version 1

In each round t, the agent

- ▶ observes a context  $c_t \in C$
- ▶ selects an arm  $A_t \in A_t$  (set of arm can vary in each round)

▶ the agent receives a reward  $r_t = f_{A_t}(c_t) + \varepsilon_t$ 

<u>Unknown</u>: regression functions  $(f_a)$  for all possible arm a

## Contextual MAB (more general)

In each round t, the agent

- ▶ is given a set of arms  $X_t$  (can be different in each round)
- ▶ selects an *arm*  $x_t \in \mathcal{X}_t$

▶ the agent receives a reward  $r_t = f_\star(x_t) + \varepsilon_t$ 

<u>Unknown</u> : regression function  $f_{\star}$ 

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<u>Unknown</u> : regression function  $f_{\star}$ 

→ **Goal** : learn the unknown function  $f_{\star}$ ... while maximizing rewards ! Emilie Kaufmann | CRIStAL

## Outline

#### **1** Contextual Bandits

### 2 Solving Linear Bandits

Lin-UCBLinear Thompson Sampling

#### 3 Other variants of the classical MAB



## **Contextual linear bandits**

In each round t, the agent

- ▶ receives a (finite) set of arms  $\mathcal{X}_t \subseteq \mathbb{R}^d$
- ▶ chooses an arm  $x_t \in \mathcal{X}_t$
- ▶ gets a reward  $r_t = \theta_{\star}^{\top} x_t + \varepsilon_t$

where

- $heta_\star \in \mathbb{R}^d$  is an unknown regression vector
- $\varepsilon_t$  is a centered noise, independent from past data

**Assumption** :  $\sigma^2$ - sub-Gaussian noise

$$\forall \lambda \in \mathbb{R}, \ \mathbb{E}\left[e^{\lambda X}\right] \leq e^{\frac{\lambda^2 \sigma^2}{2}}$$

e.g., Gaussian noise, bounded noise.

## **Contextual linear bandits**

In each round t, the agent

- $\blacktriangleright$  receives a (finite) set of arms  $\mathcal{X}_t \subseteq \mathbb{R}^d$
- ▶ chooses an arm  $x_t \in \mathcal{X}_t$
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#### (Pseudo)-regret for contextual bandit

maximizing expected total reward  $\leftrightarrow$  minimizing the expectation of

$$R_{T}(\mathcal{A}) = \sum_{t=1}^{T} \left( \max_{x \in \mathcal{X}_{t}} \theta_{\star}^{\top} x - \theta_{\star}^{\top} x_{t} \right)$$

➔ in each round, comparison to a possibly different optimal action !

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## Tools

Algorithms will rely on estimates / confidence regions / posterior distributions for  $\theta_{\star} \in \mathbb{R}^{d}$ .

• design matrix (with regularization parameter  $\lambda > 0$ )

$$B_t^{\lambda} = \lambda I_d + \sum_{s=1}^t x_s x_s^{\top}$$

regularized least-square estimate

$$\hat{\theta}_t^{\lambda} = \left(B_t^{\lambda}\right)^{-1} \left(\sum_{s=1}^t r_t x_t\right)$$

Recap from lecture 1 : easy online update !

- estimate of the expected reward of an arm  $x \in \mathbb{R}^d : x^\top \hat{\theta}_t^{\lambda}$
- → sufficient for Follow the Leader, but not for smarter algorithms !

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## How to build (tight) confidence interval on the mean rewards?

**Idea** : rely on a confidence ellippoid around  $\hat{\theta}_t^{\lambda}$ 

||x|



**Why?** For all invertible matrix positive semi-definite matrix A,

$$\begin{aligned} \forall x \in \mathbb{R}^{d}, \quad \left| x^{\top} \theta_{\star} - x^{\top} \hat{\theta}_{t}^{\lambda} \right| \leq \left\| x \right\|_{A^{-1}} \left\| \theta_{\star} - \hat{\theta}_{t}^{\lambda} \right\|_{A} \\ \| x \|_{A} = \sqrt{x^{\top} A x} \end{aligned}$$
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# How to build (tight) confidence interval on the mean rewards?

Wanted :  $\theta_{\star} \in \left\{ \theta \in \mathbb{R}^{d} : \|\theta - \hat{\theta}_{t}^{\lambda}\|_{A} \leq \beta_{t} \right\}$ 

#### Example of threshold [Abbasi-Yadkori et al., 2011]

Assuming that the noise  $\varepsilon_t$  is  $\sigma^2$ -sub-Gaussian, and that for all t and  $x \in \mathcal{X}_t$ ,  $||x|| \leq L$ , we have

$$\mathbb{P}\left(\exists t \in \mathbb{N}^{\star}: \|\theta_{\star} - \hat{\theta}^{\lambda}_{t}\|_{\boldsymbol{B}^{\lambda}_{t}} > \beta(t, \delta)\right) \leq \delta$$

with  $\beta(t, \delta) = \sigma \sqrt{2 \log (1/\delta)} + d \log \left(1 + t \frac{L}{d\lambda}\right) + \sqrt{\lambda} \|\theta_{\star}\|.$ 

→ Letting

$$C_t(\delta) = \left\{ heta \in \mathbb{R}^d : \| heta - \hat{ heta}_t^\lambda\|_{B_t^\lambda} \leq eta(t,\delta) 
ight\}.$$

one has  $\mathbb{P}(\forall t \in \mathbb{N}, \theta_{\star} \in C_t(\delta)) \geq 1 - \delta$ .

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## A Lin-UCB algorithm

#### Consequence :



$$x_{t+1} = \operatorname*{argmax}_{x \in \mathcal{X}_{t+1}} \left[ x^\top \hat{\theta}_t^{\lambda} + \|x\|_{(B_t^{\lambda})^{-1}} \beta(t, \delta) \right]$$

(many algorithms of this style, with different choices of  $\beta(t,\delta)$ ) Emilie Kaufmann | CRIStAL

## **Theoretical guarantees**

We want to bound the pseudo-regret

$$R_{T}(\text{Lin-UCB}) = \sum_{t=1}^{T} \left( \max_{x \in \mathcal{X}_{t}} \theta_{\star}^{\top} x - \theta_{\star}^{\top} x_{t} \right)$$

or its expectation, the regret  $\mathcal{R}_T(\text{Lin-UCB}) = \mathbb{E}[R_T(\text{Lin-UCB})]$ .

#### Lemma

One can prove that, with probability larger than  $1 - \delta$ ,

$$\forall T \in \mathbb{N}^*, R_T(\text{Lin-UCB}) \leq C\beta(T, \delta)\sqrt{dT\log(T)}$$

▶ with the choice of  $\beta(t, \delta)$  presented before, with high probability

$$R_T(\text{Lin-UCB}) = \mathcal{O}(d\sqrt{T}\log(T) + \sqrt{dT\log(T)\log(1/\delta)})$$

• choosing  $\delta = 1/T$ ,  $\mathcal{R}_T(\text{Lin-UCB}) = \mathcal{O}(d\sqrt{T}\log(T))$ 

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## A Bayesian view on Linear Regression

#### Bayesian model :

- ▶ likelihood :  $r_t = \theta_{\star}^{\top} x_t + \varepsilon_t$
- ▶ prior :  $\theta_{\star} \sim \mathcal{N}(0, \kappa^2 \mathsf{I}_d)$

Assuming further that the noise is Gaussian :  $\varepsilon_t \sim \mathcal{N}(0, \sigma^2)$ , the posterior distribution of  $\theta_{\star}$  has a closed form :

$$\theta_{\star}|x_{1}, r_{1}, \ldots, x_{t}, r_{t} \sim \mathcal{N}\left(\hat{\theta}_{t}^{\lambda}, \sigma^{2}\left(B_{t}^{\lambda}\right)^{-1}\right)$$

with

•  $B_t^{\lambda} = \lambda I_d + \sum_{s=1}^t x_s x_s^{\top}$ •  $\hat{\theta}_t^{\lambda} = (B_t^{\lambda})^{-1} (\sum_{s=1}^t r_s x_s)$  is the regularized least square estimate with a regularization parameter  $\lambda = \frac{\sigma^2}{\mu^2}$ .

## **Thompson Sampling for Linear Bandits**

Recall the Thompson Sampling principle :

"draw a possible model from the posterior distribution and act optimally in this sampled model"

### Thompson Sampling in linear bandits

In each round t + 1,

$$\begin{aligned} & \tilde{\theta}_t \quad \sim \quad \mathcal{N}\left(\hat{\theta}_t^{\lambda}, \sigma^2 \left(B_t^{\lambda}\right)^{-1}\right) \\ & \kappa_{t+1} \quad = \quad \operatornamewithlimits{argmax}_{x \in \mathcal{X}_{t+1}} x^\top \tilde{\theta}_t \end{aligned}$$

**Numerical complexity** : one need to draw a sample from a multivariate Gaussian distribution, e.g.

$$\tilde{\theta}_t = \hat{\theta}_t^{\lambda} + \sigma \left( B_t^{\lambda} \right)^{-1/2} X$$

where X is a vector with d independent  $\mathcal{N}(0,1)$  entries.

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## **Theoretical guarantees**

[Agrawal and Goyal, 2013] analyze a *variant* of Thompson Sampling using some "posterior inflation" :

$$\begin{aligned} \tilde{\theta}_t &\sim \mathcal{N}\left(\hat{\theta}_t^1, v^2 \left(\mathcal{B}_t^1\right)^{-1}\right) \\ \mathsf{x}_{t+1} &= \operatorname*{argmax}_{x \in \mathcal{X}_{t+1}} \mathbf{x}^\top \tilde{\theta}_t \end{aligned}$$

where  $v = \sigma \sqrt{9d \ln(T/\delta)}$ .

#### Theorem

If the noise is  $\sigma^2$ -sub-Gaussian, the above algorithm satisfies

$$\mathbb{P}\left(R_{T}(\mathrm{TS})=\mathcal{O}\left(d^{3/2}\sqrt{T}\left[\ln(T)+\sqrt{\ln(T)\ln(1/\delta)}\right]\right)\right)\geq 1-\delta.$$

slightly worse than Lin-UCB... how about in practice?

b do we need the posterior inflation?

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## **Beyond linear bandits**

Depending on the application, other parameteric models may be better suited than the simple linear model, for example the logistic model.

$$\mathbb{P}(r_t = 1|x_t) = rac{1}{1 + e^{- heta_\star^ op x_t}} \ \mathbb{P}(r_t = 0|x_t) = rac{e^{- heta_\star^ op x_t}}{1 + e^{- heta_\star^ op x_t}}$$

e.g., clic / no-clic on an add depending on a user/add feature  $x_t \in \mathbb{R}^d$ 

- [Filippi et al., 2010] : first UCB style algorithm for Generalized Linear Bandit models
- Thompson Sampling for logistic bandits [Dumitrascu et al., 2018]
- going further : UCB/TS for neural bandits !

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## Many possible structures

 $\mathcal{X}$ -armed bandits :  $\mathcal{X}_t = \mathcal{X}$  arbitrary metric space

 $r_t = f_\star(x_t) + \varepsilon_t$ 

with non-parametric assumption on  $f_{\star}$ .

Examples :

▶  $f_{\star}$  is a Lipschitz function :

$$|f_{\star}(x) - f_{\star}(y)| \leq Ld(x, y)$$

where d is a metric on  $\mathcal{X}$ .

[Bubeck et al., 2011b]

- ▶  $f_{\star}$  is a unimodal function
- $f_{\star}$  is drawn from a Gaussian process prior

[Srinivas et al., 2010]

• • • •

## Beyond one arm : Combinatorial bandits

classical bandit : one arm is selected in each round combinatorial bandit : possibility to select a group of arms (action)

e.g.,[Chen et al., 2013]



#### Example :

- arms : edges in a graph
- actions : paths from A to B
- reward : some function of the edges's rewards in the chosen path (e.g. - (total travelling distance))

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**Combinatorial bandit** : Actions  $\subseteq \mathcal{P}(\{1, \ldots, K\})$ .

In round t, the agent

- ▶ selects an action  $Act_t \in Actions$
- ▶ a reward  $r_{a,t}$  is generated for every arm  $a \in Act_t$

► the agent receives as a reward  $\sum_{a \in Act_t} r_{a,t}$  (or some other function) Emilie Kaufmann | CRIStAL

## Beyond one state : Reinforcement Learning

In most bandit models, the agent repeatedly faces the same set of actions (or at least the set of available actions in round does not depend on the past decisions).

no longer true in reinforcement learning, in which an action also triggers a transition to a new state



more on this in the next lectures

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## Bandits without rewards?



 $\mathcal{B}(\mu_1)$   $\mathcal{B}(\mu_2)$   $\mathcal{B}(\mu_3)$   $\mathcal{B}(\mu_4)$   $\mathcal{B}(\mu_5)$ 

For the *t*-th patient in a clinical study,

- chooses a treatment A<sub>t</sub>
- ▶ observes a response  $X_t \in \{0,1\}$  :  $\mathbb{P}(X_t = 1) = \mu_{A_t}$

Maximize rewards ↔ cure as many patients as possible

**Alternative goal :** identify as quickly as possible the best treatment (without trying to cure patients during the study)

## Bandits without rewards?

Probability that some version of a website generates a conversion :





**Best version** :  $a_{\star} = \underset{a=1,...,K}{\operatorname{argmax}} \mu_{a}$ 

Sequential protocol : for the *t*-th visitor :

- ▶ display version A<sub>t</sub>
- observe conversion indicator  $X_t \sim \mathcal{B}(\mu_{A_t})$ .

 $\textbf{Maximize rewards} \leftrightarrow \text{maximize the number of conversions}$ 

#### Alternative goal : identify the best version (without trying to maximize conversions during the test)

## **A Pure Exploration Problem**

**Goal :** identify an arm with mean close to  $\mu_{\star}$  as quickly and accurately as possible  $\simeq$  identify

 $a_{\star} = \underset{a=1,...,K}{\operatorname{argmax}} \mu_a.$ 

Algorithm : made of three components :

- $\rightarrow$  sampling rule :  $A_t$  (arm to explore)
- → recommendation rule :  $B_t$  (current guess for the best arm)
- $\rightarrow$  stopping rule  $\tau$  (when do we stop exploring?)

### Probability of error

The probability of error after n rounds is

 $p_{\nu}(T) = \mathbb{P}_{\nu}\left(B_T \neq a_{\star}\right).$ 

## **A Pure Exploration Problem**

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#### Simple regret [Bubeck et al., 2011a]

The simple regret after n rounds is

$$r_{\nu}(n) = \mu_{\star} - \mu_{B_n}.$$

## **A Pure Exploration Problem**

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 $a_{\star} = \underset{a=1,\ldots,K}{\operatorname{argmax}} \mu_a.$ 

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#### Simple regret [Bubeck et al., 2011a]

The simple regret after n rounds is

$$r_{\nu}(n)=\mu_{\star}-\mu_{B_n}.$$

$$\Delta_{\min} p_{
u}(\mathcal{T}) \leq \mathbb{E}_{
u}[r_{
u}(\mathcal{T})] \leq \Delta_{\max} p_{
u}(\mathcal{T})$$

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## **Several objectives**

Algorithm : made of three components :

- $\rightarrow$  sampling rule :  $A_t$  (arm to explore)
- → recommendation rule :  $B_t$  (current guess for the best arm)
- $\rightarrow$  stopping rule  $\tau$  (when do we stop exploring?)
- ► Objectives studied in the literature :

Fixed-budget setting	Fixed-confidence setting
input : budget T	input : risk parameter $\delta$
	(tolerance parameter $\epsilon$ )
au = T	minimize $\mathbb{E}[ au]$
minimize $\mathbb{P}(B_{\mathcal{T}}  eq a_{\star})$	$\mathbb{P}(B_ au  eq a_\star) \leq \delta$
or $\mathbb{E}[r_{\mathcal{T}}( u)]$	or $\mathbb{P}(r_{ u}( au) > \epsilon) \leq \delta$
[Bubeck et al., 2011a]	[Even-Dar et al., 2006]
[Audibert et al., 2010]	

**Context** : bounded rewards ( $\nu_a$  supported in [0, 1]) We know good algorithms to maximize rewards, for example UCB( $\alpha$ )

$$A_{t+1} = \underset{a=1,...,\mathcal{K}}{\operatorname{argmax}} \hat{\mu}_{a}(t) + \sqrt{\frac{\alpha \ln(t)}{N_{a}(t)}}$$

▶ How good is it for best arm identification?

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▶ How good is it for best arm identification ?

Possible recommendation rules :

Empirical Best Arm	$B_t = \operatorname{argmax}_a \hat{\mu}_a(t)$
(EBA)	
Most Played Arm	$B_t = \operatorname{argmax}_a N_a(t)$
(MPA)	
Empirical Distribution of Plays	$B_t \sim p_t$ , where
(EDP)	$p_t = \left(\frac{N_1(t)}{t}, \ldots, \frac{N_{\kappa}(t)}{t}\right)$

[Bubeck et al., 2011a]

**Context** : bounded rewards ( $\nu_a$  supported in [0, 1]) We know good algorithms to maximize rewards, for example UCB( $\alpha$ )

$$\mathcal{A}_{t+1} = \operatorname*{argmax}_{a=1,...,\mathcal{K}} \hat{\mu}_{a}(t) + \sqrt{rac{lpha \ln(t)}{N_{a}(t)}}$$

▶ How good is it for best arm identification ?

Possible recommendation rules :

Empirical Best Arm (EBA)	$B_t = \operatorname{argmax}_a \hat{\mu}_a(t)$
Most Played Arm (MPA)	$B_t = \operatorname{argmax}_a N_a(t)$
Empirical Distribution of Plays	$B_t \sim p_t$ , where
(EDP)	$p_t = \left(\frac{N_1(t)}{t}, \dots, \frac{N_K(t)}{t}\right)$

[Bubeck et al., 2011a]

UCB + Empirical Distribution of Plays

$$\mathbb{E}[r_{\nu}(T)] = \mathbb{E}\left[\mu_{\star} - \mu_{B_{T}}\right] = \mathbb{E}\left[\sum_{b=1}^{K} (\mu_{\star} - \mu_{b})\mathbb{1}_{(B_{T}=b)}\right]$$
$$= \mathbb{E}\left[\sum_{b=1}^{K} (\mu_{\star} - \mu_{b})\mathbb{P}(B_{T}=b|\mathcal{F}_{T})\right]$$
$$= \mathbb{E}\left[\sum_{b=1}^{K} (\mu_{\star} - \mu_{b})\frac{N_{b}(T)}{T}\right]$$
$$= \frac{1}{T}\sum_{b=1}^{K} (\mu_{\star} - \mu_{b})\mathbb{E}[N_{b}(T)]$$
$$= \frac{\mathcal{R}_{\nu}(T)}{T}.$$

→ a conversion from cumulative regret to simple regret !

#### UCB + Empirical Distribution of Plays

$$\mathbb{E}\left[r_{\nu}\left(\texttt{UCB}(\alpha), T\right)\right] \leq \frac{\mathcal{R}_{\nu}(\texttt{UCB}(\alpha), T)}{T} \leq \frac{\mathcal{C}(\nu) \ln(T)}{T}$$

#### UCB + Empirical Distribution of Plays

$$\mathbb{E}\left[r_{\nu}\left(\mathtt{UCB}(\alpha), T\right)\right] \leq \frac{\mathcal{R}_{\nu}(\mathtt{UCB}(\alpha), T)}{T} \leq C\sqrt{\frac{K\ln(T)}{T}}$$

#### UCB + Empirical Distribution of Plays

$$\mathbb{E}\left[r_{\nu}\left(\mathtt{UCB}(\alpha), T\right)\right] \leq \frac{\mathcal{R}_{\nu}(\mathtt{UCB}(\alpha), T)}{T} \leq C\sqrt{\frac{K\ln(T)}{T}}$$

#### vs. Uniform Sampling

The simple regret or the uniform strategy decays exponentially :

$$\mathbb{E}_{
u}\left[ \textit{r}_{
u}\left(\texttt{Unif}, T
ight) 
ight] \leq (\mathcal{K}-1) \Delta_{\max} \exp\left(-rac{1}{2}rac{T}{\mathcal{K}}\Delta_{\min}^2
ight)$$

→ UCB does not provably outperform uniform sampling...

## Fixed Budget : Sequential Halving

#### **Input** : total number of plays T

**Idea :** split the budget in  $\log_2(K)$  phases of equal length, eliminate the worst half of the remaining arms after each phase.

**Initialisation** : 
$$S_0 = \{1, ..., K\}$$
;  
**For**  $r = 0$  **to**  $\lceil \ln_2(K) \rceil - 1$ , **do**  
sample each arm  $a \in S_r$   $t_r = \lfloor \frac{T}{|S_r| \lceil \log_2(K) \rceil} \rfloor$  times;  
let  $\hat{\mu}_a^r$  be the empirical mean of arm  $a$ ;  
let  $S_{r+1}$  be the set of  $\lceil |S_r|/2 \rceil$  arms with largest  $\hat{\mu}_a^r$   
**Output** :  $B_T$  the unique arm in  $S_{\lceil \log_2(K) \rceil}$ 

### Theorem [Karnin et al., 2013]

Letting 
$$H_2(\nu) = \max_{a \neq a_\star} a \Delta_{[a]}^{-2}$$
, for any bounded bandit instance,  
 $\mathbb{P}_{\nu} \left( B_T \neq a_\star \right) \le 3 \log_2(K) \exp\left(-\frac{T}{8 \log_2(K) H_2(\nu)}\right).$ 

## Fixed Budget : LUCB

 $\mathcal{I}_{a}(t) = [LCB_{a}(t), UCB_{a}(t)].$ 



### Theorem [Kalyanakrishnan et al., 2012]

For well-chosen confidence intervals,  $\mathbb{P}_{\nu}(\mu_{B_{\tau}} > \mu_{\star} - \epsilon) \geq 1 - \delta$  and

$$\mathbb{E}\left[\tau_{\delta}\right] = O\left(\left[\frac{1}{\Delta_2^2 \vee \epsilon^2} + \sum_{\mathsf{a}=2}^{\mathsf{K}} \frac{1}{\Delta_{\mathsf{a}}^2 \vee \epsilon^2}\right] \ln\left(\frac{1}{\delta}\right)\right)$$

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# (kl)-LUCB in action

$$\begin{split} &\text{UCB}_{\mathsf{a}}(t) = \max \left\{ q \in [0,1] : N_{\mathsf{a}}(t) \text{kl}(\hat{\mu}_{\mathsf{a}}(t),q) \leq \log(Ct^2/\delta) \right\} \\ &\text{LCB}_{\mathsf{a}}(t) = \min \left\{ q \in [0,1] : N_{\mathsf{a}}(t) \text{kl}(\hat{\mu}_{\mathsf{a}}(t),q) \leq \log(Ct^2/\delta) \right\} \end{split}$$



## A comparison with UCB

Regret minimizing algorithms and Best Arm Identification algorithms behave quite differently



Number of selections and confidence intervals for KL-UCB (left) and KL-LUCB (right)

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