Reinforcement Learning Multi-Armed Bandits

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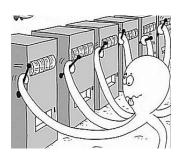


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Stochastic bandit: a simple MDP

A stochastic multi-armed bandit model can be viewed as an MDP with a single state s_0

- ▶ unknown reward distribution $\nu_{s_0,a}$ with mean $r(s_0,a)$
- ightharpoonup transition $p(s_0|s_0,a)=1$
- ▶ the agent repeatedly chooses between the same set of actions



an agent facing arms in a Multi-Armed Bandit

Sequential resource allocation

Clinical trials

▶ K treatments for a given symptom (with unknown effect)













► What treatment should be allocated to the next patient based on responses observed on previous patients?

Online advertisement

K adds that can be displayed









▶ Which add should be displayed for a user, based on the previous clicks of previous (similar) users?

The Multi-Armed Bandit Setup

K arms $\leftrightarrow K$ rewards streams $(X_{a,t})_{t\in\mathbb{N}}$











At round t, an agent :

- \triangleright chooses an arm A_t
- ightharpoonup receives a reward $R_t = X_{A_t,t}$

Sequential sampling strategy (bandit algorithm) :

$$A_{t+1} = F_t(A_1, R_1, \ldots, A_t, R_t).$$

Goal: Maximize $\sum_{t=1}^{T} R_t$.

The Stochastic Multi-Armed Bandit Setup

K arms \leftrightarrow K probability distributions : ν_a has mean μ_a











 ν_{5}

At round t, an agent :

- \triangleright chooses an arm A_t
- ightharpoonup receives a reward $R_t = X_{A_t,t} \sim \nu_{A_t}$

Sequential sampling strategy (bandit algorithm) :

$$A_{t+1} = F_t(A_1, R_1, \ldots, A_t, R_t).$$

Goal : Maximize $\mathbb{E}\left[\sum_{t=1}^{T} R_{t}\right]$

→ a particular reinforcement learning problem

Clinical trials

Historical motivation [Thompson, 1933]



For the t-th patient in a clinical study,

- chooses a treatment A_t
- $lackbox{ observes a response } R_t \in \{0,1\}: \mathbb{P}(R_t=1|A_t=a)=\mu_{a}$

Goal: maximize the expected number of patients healed

Online content optimization

Modern motivation (\$\$) [Li et al., 2010] (recommender systems, online advertisement)











For the t-th visitor of a website,

- ightharpoonup recommend a movie A_t
- ▶ observe a rating $R_t \sim \nu_{A_t}$ (e.g. $R_t \in \{1, ..., 5\}$)

Goal: maximize the sum of ratings

Outline

1 Performance measure and first strategies

Mixing Exploration and ExploitationUpper Confidence Bound algorithms

- 3 Bayesian bandit algorithms

 Thompson Sampling
 - Thompson Sampling

Regret of a bandit algorithm

Bandit instance : $\nu = (\nu_1, \nu_2, \dots, \nu_K)$, mean of arm $a : \mu_a = \mathbb{E}_{X \sim \nu_a}[X]$.

$$\mu_{\star} = \max_{a \in \{1, \dots, K\}} \mu_a$$
 $a_{\star} = \operatorname*{argmax}_{a \in \{1, \dots, K\}} \mu_a$.

Maximizing rewards \leftrightarrow selecting a_{\star} as much as possible \leftrightarrow minimizing the regret [Robbins, 1952]

$$\mathcal{R}_{
u}(\mathcal{A}, \mathcal{T}) := \underbrace{\mathcal{T}\mu_{\star}}_{\substack{\text{sum of rewards of an oracle strategy always selecting } a_{\star}}} - \underbrace{\mathbb{E}\left[\sum_{t=1}^{\mathcal{T}} R_{t}\right]}_{\substack{\text{sum of rewards of the strategy} \mathcal{A}}}$$

What regret rate can we achieve?

- ightharpoonup consistency : $\frac{\mathcal{R}_{\nu}(\mathcal{A}, \mathcal{T})}{\mathcal{T}}
 ightarrow 0$
- → can we be more precise?

Regret decomposition

 $N_a(t)$: number of selections of arm a in the first t rounds $\Delta_a:=\mu_\star-\mu_a$: sub-optimality gap of arm a

Regret decomposition

$$\mathcal{R}_{\nu}(\mathcal{A},T) = \sum_{a=1}^{K} \Delta_{a} \mathbb{E}\left[N_{a}(T)\right].$$

Proof.



Regret decomposition

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Regret decomposition

$$\mathcal{R}_{\nu}(\mathcal{A},T) = \sum_{a=1}^{K} \Delta_{a} \mathbb{E}\left[N_{a}(T)\right].$$

A strategy with small regret should :

- ▶ select not too often arms for which $\Delta_a > 0$
- \blacktriangleright ... which requires to try all arms to estimate the values of the Δ_a 's

⇒ Exploration / Exploitation trade-off

Two naive strategies

▶ Idea 1 : Uniform Exploration

Draw each arm T/K times

$$\Rightarrow$$
 EXPLORATION $\mathcal{R}_{\nu}(\mathcal{A},T) = \left(\frac{1}{K}\sum_{a:u_a>u_a}\Delta_a\right)T$

Two naive strategies

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u}(\mathcal{A},T) = \left(\frac{1}{K}\sum_{a:\mu_a>\mu_\star}\Delta_a\right)T$

▶ Idea 2 : Follow The Leader

$$\begin{aligned} A_{t+1} &= \underset{a \in \{1, \dots, K\}}{\operatorname{argmax}} \ \hat{\mu}_a(t) \\ \hat{\mu}_a(t) &= \frac{1}{N_a(t)} \sum_{s=1}^t X_{a,s} \mathbb{1}_{(A_s = a)} \end{aligned}$$

is an estimate of the unknown mean μ_a .

$$\Rightarrow$$
 EXPLOITATION $\mathcal{R}_{\nu}(\mathcal{A}, T) \geq (1 - \mu_1) \times \mu_2 \times (\mu_1 - \mu_2) T$ (Bernoulli arms)

Given $m \in \{1, \ldots, T/K\}$,

- draw each arm m times
- ▶ compute the empirical best arm $\hat{a} = \operatorname{argmax}_a \hat{\mu}_a(Km)$
- ▶ keep playing this arm until round *T*

$$A_{t+1} = \hat{a}$$
 for $t \geq Km$

⇒ EXPLORATION followed by EXPLOITATION

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⇒ EXPLORATION followed by EXPLOITATION

Analysis for two arms. $\mu_1 > \mu_2$, $\Delta := \mu_1 - \mu_2$.

$$\begin{array}{lcl} \mathcal{R}_{\nu}(\texttt{ETC},\,\mathcal{T}) & = & \Delta \mathbb{E}[\textit{N}_{2}(\textit{T})] \\ & = & \Delta \mathbb{E}\left[\textit{m} + (\textit{T} - 2\textit{m})\mathbb{1}\left(\hat{\textit{a}} = 2\right)\right] \\ & \leq & \Delta \textit{m} + (\Delta \textit{T}) \times \mathbb{P}\left(\hat{\mu}_{2,\textit{m}} \geq \hat{\mu}_{1,\textit{m}}\right) \end{array}$$

 $\hat{\mu}_{a,m}$: empirical mean of the first m observations from arm a

Given $m \in \{1, \ldots, T/K\}$,

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Analysis for two arms. $\mu_1 > \mu_2$, $\Delta := \mu_1 - \mu_2$.

$$\mathcal{R}_{\nu}(\text{ETC}, T) = \Delta \mathbb{E}[N_2(T)]$$

$$= \Delta \mathbb{E}[m + (T - 2m)\mathbb{1}(\hat{a} = 2)]$$

$$\leq \Delta m + (\Delta T) \times \mathbb{P}(\hat{\mu}_{2,m} \geq \hat{\mu}_{1,m})$$

 $\hat{\mu}_{a,m}$: empirical mean of the first m observations from arm $a \to \text{requires a concentration inequality}$

Intermezzo: Concentration Inequalities

Sub-Gaussian random variables : Z is σ^2 -subGaussian if

$$\mathbb{E}[Z] = \mu \text{ and } \mathbb{E}\left[e^{\lambda(Z-\mu)}\right] \le e^{\frac{\lambda^2\sigma^2}{2}}.$$
 (1)

Hoeffding inequality

 Z_i i.i.d. satisfying (1). For all $s \geq 1$

$$\mathbb{P}\left(\frac{Z_1 + \dots + Z_s}{s} \ge \mu + x\right) \le e^{-\frac{sx^2}{2\sigma^2}}$$

Proof: Cramér-Chernoff method

- $\triangleright \nu_a$ bounded in $[a, b] : (b a)^2/4$ sub-Gaussian (Hoeffding's lemma)
- $\nu_a = \mathcal{N}(\mu_a, \sigma^2)$: σ^2 sub-Gaussian

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Analysis for two arms. $\mu_1 > \mu_2$, $\Delta := \mu_1 - \mu_2$.

Assumption : ν_1, ν_2 are bounded in [0, 1].

$$\mathcal{R}_{\nu}(T) = \Delta \mathbb{E}[N_2(T)]$$

$$= \Delta \mathbb{E}[m + (T - 2m)\mathbb{1}(\hat{a} = 2)]$$

$$< \Delta m + (\Delta T) \times \mathbb{P}(\hat{\mu}_{2m} > \hat{\mu}_{1m})$$

 $\hat{\mu}_{a,m}$: empirical mean of the first m observations from arm $a \to \mathsf{Hoeffding's}$ inequality

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$$< \Delta m + (\Delta T) \times \exp(-m\Delta^2/2)$$

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Assumption: ν_1, ν_2 are bounded in [0, 1].

For
$$m = \frac{2}{\Delta^2} \log \left(\frac{T\Delta^2}{2} \right)$$
,

$$\mathcal{R}_{
u}(exttt{ETC}, \mathcal{T}) \leq rac{2}{\Delta} \left\lceil \log \left(rac{\mathcal{T}\Delta^2}{2}
ight) + 1
ight
ceil.$$

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ight) + 1
ight].$$

- + logarithmic regret!
- requires the knowledge of T and Δ

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A simple strategy : ϵ -greedy

The ϵ -greedy rule [Sutton and Barto, 1998] is the simplest way to alternate exploration and exploitation.

ϵ -greedy strategy

At round *t*,

ightharpoonup with probability ϵ

$$A_t \sim \mathcal{U}(\{1,\ldots,K\})$$

ightharpoonup with probability $1-\epsilon$

$$A_t = \underset{a=1,...,K}{\operatorname{argmax}} \hat{\mu}_a(t).$$

→ Linear regret : \mathcal{R}_{ν} (ϵ -greedy, T) $\geq \epsilon \frac{K-1}{K} \Delta_{\min} T$.

$$\Delta_{\min} = \min_{a:u_a < u_a} \Delta_a$$

A simple strategy : ϵ -greedy

A simple fix:

ϵ_t -greedy strategy

At round t,

• with probability $\epsilon_t := \min \left(1, \frac{K}{d^2 t}\right)$

$$A_t \sim \mathcal{U}(\{1,\ldots,K\})$$

 \blacktriangleright with probability $1 - \epsilon_t$

$$A_t = \underset{a=1,...,K}{\operatorname{argmax}} \hat{\mu}_a(t-1).$$

Theorem [Auer, 2002]

If
$$0 < d \leq \Delta_{\min}$$
, $\mathcal{R}_{
u}\left(\epsilon_t ext{-greedy}, T\right) = O\left(rac{K\log(T)}{d^2}
ight)$.

 \rightarrow requires the knowledge of a lower bound on Δ_{\min} ...

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The optimism principle

Step 1: construct a set of statistically plausible models

▶ For each arm a, build a confidence interval on the mean μ_a :

$$\mathcal{I}_{a}(t) = [\mathrm{LCB}_{a}(t), \mathrm{UCB}_{a}(t)]$$

LCB = Lower Confidence Bound UCB = Upper Confidence Bound

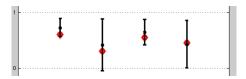


FIGURE - Confidence intervals on the means after t rounds

The optimism principle

Step 2: act as if the best possible model were the true model (optimism in face of uncertainty)

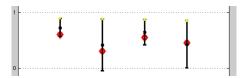


FIGURE – Confidence intervals on the means after t rounds

Optimistic bandit model =
$$\operatorname*{argmax}_{\boldsymbol{\mu} \in \mathcal{C}(t)} \operatorname*{max}_{\boldsymbol{\sigma} = 1, \dots, K} \boldsymbol{\mu}_{\boldsymbol{\sigma}}$$

▶ That is, select

$$A_{t+1} = \underset{a=1,\dots,K}{\operatorname{argmax}} \operatorname{UCB}_a(t).$$

We need $UCB_a(t)$ such that

$$\mathbb{P}\left(\mu_{\mathsf{a}} \leq \mathrm{UCB}_{\mathsf{a}}(t)\right) \gtrsim 1 - t^{-1}.$$

→ tool : concentration inequalities

Example : rewards are σ^2 sub-Gaussian

Hoeffding inequality, reloaded

 Z_i i.i.d. satisfying (1). For all $s \geq 1$

$$\mathbb{P}\left(\frac{Z_1 + \dots + Z_s}{s} < \mu - x\right) \le e^{-\frac{sx^2}{2\sigma^2}}$$

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Cannot be used directly in a bandit model as the number of observations from each arm is random!

- $N_a(t) = \sum_{s=1}^t \mathbb{1}_{(A_s=a)}$ number of selections of a after t rounds
- $\hat{\mu}_{a,s} = \frac{1}{s} \sum_{k=1}^{s} Y_{a,k}$ average of the first s observations from arm a
- $\hat{\mu}_a(t) = \hat{\mu}_{a,N_a(t)}$ empirical estimate of μ_a after t rounds

Hoeffding inequality + union bound

$$\mathbb{P}\left(\mu_{\mathsf{a}} \leq \hat{\mu}_{\mathsf{a}}(t) + \sigma \sqrt{\frac{\beta \log(t)}{N_{\mathsf{a}}(t)}}\right) \geq 1 - \frac{1}{t^{\frac{\beta}{2} - 1}}$$

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Proof.

$$\mathbb{P}\left(\mu_{a} > \hat{\mu}_{a}(t) + \sigma\sqrt{\frac{\beta \log(t)}{N_{a}(t)}}\right) \leq \mathbb{P}\left(\exists s \leq t : \mu_{a} > \hat{\mu}_{a,s} + \sigma\sqrt{\frac{\beta \log(t)}{s}}\right)$$
$$\leq \sum_{s=1}^{t} \mathbb{P}\left(\hat{\mu}_{a,s} < \mu_{a} - \sigma\sqrt{\frac{\beta \log(t)}{s}}\right) \leq \sum_{s=1}^{t} \frac{1}{t^{\beta/2}} = \frac{1}{t^{\beta/2-1}}.$$

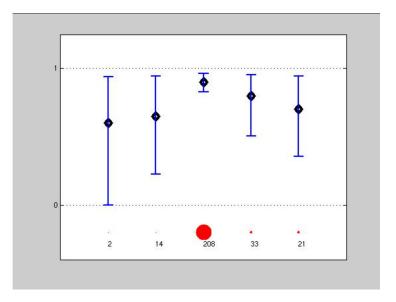
A first UCB algorithm

 $UCB(\alpha)$ selects $A_{t+1} = \operatorname{argmax}_a \ UCB_a(t)$ where

$$\mathrm{UCB}_{a}(t) = \underbrace{\hat{\mu}_{a}(t)}_{\text{exploitation term}} + \underbrace{\sqrt{\frac{\alpha \log(t)}{N_{a}(t)}}}_{\text{exploration bonus}}.$$

- ▶ popularized by [Auer, 2002] for bounded rewards : UCB1, for $\alpha = 2$
- ▶ the analysis was UCB(α) was further refined to hold for $\alpha > 1/2$, still for bounded rewards [Bubeck, 2010]

A UCB algorithm in action



Regret of UCB(α)

Context : σ^2 sub-Gaussian rewards

$$UCB_a(t) = \hat{\mu}_a(t) + \sqrt{\frac{2\sigma^2(\log(t) + c\log\log(t))}{N_a(t)}}$$

Theorem [Cappé et al.'13]

For $c \ge 3$, the UCB algorithm associated to the above index satisfy

$$\mathbb{E}[N_a(T)] \leq \frac{2\sigma^2}{(\mu_{\star} - \mu_a)^2} \log(T) + C_{\mu} \sqrt{\log(T)}.$$

if the rewards distributions are σ^2 sub-Gaussian.

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if the rewards distributions are σ^2 sub-Gaussian.

 \blacktriangleright regret bound for Gaussian distribution with variance σ^2 :

$$\mathcal{R}_{\nu}(\mathrm{UCB}(\alpha), T) = 2\sigma^{2} \left(\sum_{\mathbf{a}: \mu_{\mathbf{a}} < \mu_{\star}} \frac{1}{\Delta_{\mathbf{a}}} \right) \log(T) + \mathcal{O}(\sqrt{\log(T)})$$

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if the rewards distributions are σ^2 sub-Gaussian.

▶ regret bound for distributions that are bounded in [0,1]:

$$\mathcal{R}_{
u}(\mathrm{UCB}(lpha),\,\mathcal{T}) = rac{1}{2} \left(\sum_{a:u_{\star} < u_{\star}} rac{1}{\Delta_{a}}
ight) \log(\mathcal{T}) + \mathcal{O}(\sqrt{\log(\mathcal{T})})$$

for
$$\alpha = 1/2$$
.

Is $UCB(\alpha)$ the best possible algorithm?

Context: a parametric bandit model where each arm is parameterized by its mean $\nu = (\nu_{\mu_1}, \dots, \nu_{\mu_K}), \ \mu_a \in \mathcal{I}.$

$$\nu \leftrightarrow \boldsymbol{\mu} = (\mu_1, \dots, \mu_K)$$

Key tool: Kullback-Leibler divergence.

Kullback-Leibler divergence

$$\mathrm{kl}(\mu,\mu') := \mathsf{KL}\left(
u_{\mu},
u_{\mu'}
ight) = \mathbb{E}_{X \sim
u_{\mu}}\left[\lograc{d
u_{\mu}}{d
u_{\mu'}}(X)
ight]$$

Lower bound [Lai and Robbins, 1985]

For uniformly good algorithm,

$$\mu_{\mathsf{a}} < \mu_{\star} \Rightarrow \liminf_{T o \infty} \frac{\mathbb{E}_{\mu}[\mathsf{N}_{\mathsf{a}}(T)]}{\log T} \geq \frac{1}{\mathrm{kl}(\mu_{\mathsf{a}}, \mu_{\star})}$$

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Kullback-Leibler divergence

$$kl(\mu, \mu') := \frac{(\mu - \mu')^2}{2\sigma^2}$$
 (Gaussian bandits)

Lower bound [Lai and Robbins, 1985]

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Kullback-Leibler divergence

$$\mathrm{kl}(\mu,\mu') := \mu \log \left(\frac{\mu}{\mu'}\right) + (1-\mu) \log \left(\frac{1-\mu}{1-\mu'}\right) \quad \text{(Bernoulli bandits)}$$

Lower bound [Lai and Robbins, 1985]

For uniformly good algorithm,

$$\mu_{\mathsf{a}} < \mu_{\star} \Rightarrow \liminf_{T \to \infty} \frac{\mathbb{E}_{\mu}[N_{\mathsf{a}}(T)]}{\log T} \geq \frac{1}{\mathrm{kl}(\mu_{\mathsf{a}}, \mu_{\star})}$$

For Gaussian bandits with variance σ^2 .

▶ Upper bound for UCB($2\sigma^2$) :

$$\mathcal{R}_{\nu}(\text{UCB}, T) \lesssim \sum_{a: \mu_a < \mu_{\star}} \frac{2\sigma^2}{(\mu^{\star} - \mu_a)} \log(T)$$

▶ Lower bound : for large values of T,

$$\mathcal{R}_{\nu}(\mathcal{A}, T) \gtrsim \sum_{a: \mu_a < \mu_{\star}} \frac{(\mu_{\star} - \mu_a)}{\mathrm{kl}(\mu_a, \mu_{\star})} \log{(T)}$$

For Gaussian bandits with variance σ^2 ,

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→ UCB is asymptotically optimal for Gaussian bandits!

For Bernoulli bandits (that are bounded in [0,1]),

▶ Upper bound for UCB(1/2):

$$\mathcal{R}_{\nu}(\mathrm{UCB}, T) \lesssim \sum_{a: \mu_a < \mu_\star} \frac{1}{2(\mu^\star - \mu_a)} \log(T)$$

▶ Lower bound : for large values of T,

$$\mathcal{R}_{\nu}(\mathcal{A}, T) \gtrsim \sum_{a: \mu_a < \mu_{\star}} \frac{(\mu_{\star} - \mu_a)}{\mathrm{kl}(\mu_a, \mu_{\star})} \log(T)$$

For Bernoulli bandits (that are bounded in [0,1]),

▶ Upper bound for UCB(1/2) :

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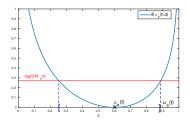
→ UCB is *not* asymptotically optimal for Bernoulli bandits...

Pinsker's inequality : $kl(\mu, \mu') \ge 2(\mu - \mu')^2$

The kl-UCB algorithm

Exploits the KL-divergence in the lower bound!

$$\mathrm{UCB}_{\mathsf{a}}(t) = \max \left\{ q \in [0,1] : \mathrm{kl}\left(\hat{\mu}_{\mathsf{a}}(t),q\right) \leq \frac{\log(t)}{\mathcal{N}_{\mathsf{a}}(t)}
ight\}.$$



A tighter concentration inequality [Garivier and Cappé, 2011]

For Bernoulli rewards

$$\mathbb{P}(\mathrm{UCB}_{\mathsf{a}}(t) > \mu_{\mathsf{a}}) \gtrsim 1 - \frac{1}{t \log(t)}.$$

An asymptotically optimal algorithm

kl-UCB selects $A_{t+1} = \operatorname{argmax}_{a} \operatorname{UCB}_{a}(t)$ with

$$\mathrm{UCB}_{\mathsf{a}}(t) = \max \left\{ q \in [0,1] : \mathrm{kl}\left(\hat{\mu}_{\mathsf{a}}(t),q\right) \leq \frac{\log(t) + c \log\log(t)}{\mathcal{N}_{\mathsf{a}}(t)} \right\}.$$

Theorem [Cappé et al., 2013]

If $c \geq 3$, for every arm such that $\mu_a < \mu_{\star}$,

$$\mathbb{E}_{\mu}[N_{a}(T)] \leq \frac{1}{\mathrm{kl}(\mu_{a}, \mu_{\star})} \log(T) + C_{\mu} \sqrt{\log(T)}.$$

▶ kl-UCB is asymptotically optimal for Bernoulli bandits :

$$\mathcal{R}_{m{\mu}}(ext{kl-UCB}, T) \simeq \left(\sum_{m{a}: \mu_\star < \mu_\star} rac{\mu_\star - \mu_ ext{a}}{ ext{kl}(\mu_ ext{a}, \mu_\star)}
ight) \log(T).$$

Outline

1 Performance measure and first strategies

Mixing Exploration and ExploitationUpper Confidence Bound algorithms

- 3 Bayesian bandit algorithms
 - Thompson Sampling

Frequentist versus Bayesian bandit

Context: parametric bandit model $\nu_{\mu} = (\nu_{\mu_1}, \dots, \nu_{\mu_K})$.

► Two probabilistic models

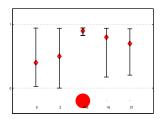
Frequentist model	Bayesian model
μ_1,\ldots,μ_K	μ_1,\ldots,μ_K drawn from a
unknown parameters	prior distribution : $\mu_{\sf a} \sim \pi_{\sf a}$
arm $a: (Y_{a,s})_s \overset{i.i.d.}{\sim} u_{\mu_{a}}$	arm $a:(Y_{a,s})_s \mu\stackrel{\text{i.i.d.}}{\sim} \nu_{\mu_a}$

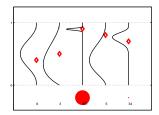
where $(Y_{a,s})$ is the sequence of successive rewards obtained from arm a

Frequentist and Bayesian algorithms

▶ Two types of tools to build bandit algorithms :

Frequentist tools	Bayesian tools
MLE estimators of the means Confidence Intervals	Posterior distributions $\pi_a^t = \mathcal{L}(\mu_a Y_{a,1},\ldots,Y_{a,N_a(t)})$





Example: Bernoulli bandits

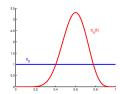
Bernoulli bandit model $\mu = (\mu_1, \dots, \mu_K)$

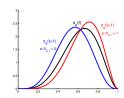
- **Bayesian view** : μ_1, \dots, μ_K are random variables prior distribution : $\mu_a \sim \mathcal{U}([0,1])$
- posterior distribution :

$$\pi_{a}(t) = \mathcal{L}(\mu_{a}|R_{1}, \dots, R_{t})$$

$$= \operatorname{Beta}\left(\underbrace{S_{a}(t)}_{\#ones} + 1, \underbrace{N_{a}(t) - S_{a}(t)}_{\#zeros} + 1\right)$$

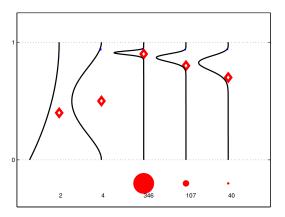
 $S_a(t) = \sum_{s=1}^t R_s \mathbb{1}_{(A_s=a)}$ sum of the rewards.





Bayesian algorithm

A Bayesian bandit algorithm exploits the posterior distributions of the means to decide which arm to select.



Outline

1 Performance measure and first strategies

- 2 Mixing Exploration and Exploitation
 - Upper Confidence Bound algorithms

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Thompson Sampling

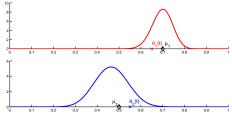
A very old idea: [Thompson, 1933].

Two equivalent interpretations:

- "select an arm at random according to its probability of being the best"

Thompson Sampling: a randomized Bayesian algorithm

$$\left\{ \begin{array}{l} \forall a \in \{1..K\}, \quad \theta_a(t) \sim \pi_a(t) \\ A_{t+1} = \mathop{\mathsf{argmax}}_{a=1...K} \theta_a(t). \end{array} \right.$$



Thompson Sampling is asymptotically optimal

Problem-dependent regret

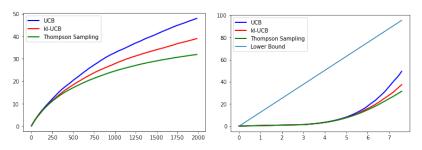
$$\forall \epsilon > 0, \quad \mathbb{E}_{\mu}[N_{a}(T)] \leq (1+\epsilon) \frac{1}{\mathrm{kl}(\mu_{a}, \mu_{\star})} \log(T) + o_{\mu, \epsilon}(\log(T)).$$

This results holds:

- ► for Bernoulli bandits, with a uniform prior [Kaufmann et al., 2012, Agrawal and Goyal, 2013]
- ▶ for Gaussian bandits, with Gaussian prior [Agrawal and Goyal, 2017]
- ► for exponential family bandits, with Jeffrey's prior [Korda et al., 2013]

Bayesian versus Frequentist algorithms

Regret up to T = 2000 (average over N = 200 runs) as a function of T (resp. log(T))



$$\mu = [0.1 \ 0.15 \ 0.2 \ 0.25]$$

Summary

Several ways to solve the exploration/exploitation trade-off, mostly

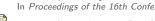
- ▶ the optimism-in-face-of-uncertainty principle (UCB)
- posterior sampling (Thompson Sampling)

What do they need?

- ▶ UCB : the capacity to build a confidence region for the unknown model parameters and compute the best possible model
- ► Thompson Sampling : the ability to define a prior distribution and sample from the corresponding posterior distribution
- → these principles can be extended to more challenging bandit problems and to reinforcement learning



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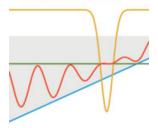


On the likelihood that one unknown probability exceeds another in view of the evidence of two samples.



Bandit Algorithms

TOR LATTIMORE CSABA SZEPESVÁRI



The Bandit Book

by [Lattimore and Szepesvari, 2019]